

A penalized maximum likelihood estimator for the segmentation of RNA-Seq data

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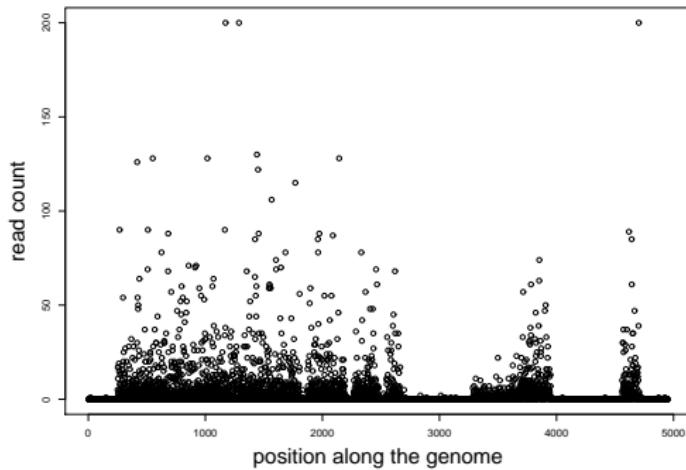
Outline

1 Motivation and main result

2 Scheme of the proof

3 Illustration

Segmentation Model



- m a partition of $\{1, n\}$,
- J a segment of m ,
- K the number of segments,

$$\forall t \in J, Y_t \sim \mathcal{NB}(p_J, \phi)$$

Penalized log-likelihood framework

$$s(t) = \mathcal{N}\mathcal{B}(p_t, \phi) \quad \text{the true model}$$

Collection of models $\mathcal{S}_m = \{s_m \mid \forall J \in m, \forall t \in J, s_m(t) = \mathcal{N}\mathcal{B}(p_J, \phi)\}.$

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Minimal contrast estimator on partition m : $\hat{s}_m = \arg \min_{u \in \mathcal{S}_m} \gamma(u)$

Log-likelihood contrast $\gamma(u) = \sum_{t=1}^n -\phi \log p_t - Y_t \log(1 - p_t),$
 \rightarrow defines collection $\mathcal{S} = \{(\hat{s}_m)_{m \in \mathcal{M}_n}\}$

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→ defines collection $\mathcal{S} = \{(\hat{s}_m)_{m \in \mathcal{M}_n}\}$

→ Choose best estimator $\hat{s}_{m(s)} = \arg \min_{u \in \mathcal{S}} \mathbf{E}[K(s, u)]$

→ Requires the knowledge of s .

Penalized log-likelihood framework (2)

Penalized likelihood estimator $\hat{s}_{\hat{m}}$ such that:

$$\hat{m} = \arg \min_{m \in \mathcal{M}_n} \{\gamma(\hat{s}_m) + \text{pen}(m)\},$$

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Oracle inequality

$$\mathbf{E}[K(s, \hat{s}_{\hat{m}})] \leq C \mathbf{E}[K(s, \hat{s}_{m(s)})],$$

Our proposition

Assumptions:

- $\forall t, 0 < \rho_{min} \leq p_t \leq \rho_{max} < 1$ and
- $\forall J \in m_f, |J| \geq \Gamma(\log(n))^2$.

Let $\beta > 1/4$ and

$$pen(m) = \beta|m| \left(1 + 4\sqrt{1.1 + \log\left(\frac{n}{|m|}\right)} \right)^2, \quad \text{then}$$

$$\mathbf{E} [h^2(s, \hat{s}_{\hat{m}})] \leq C \log(n) \inf_{m \in \mathcal{M}_n} \{\mathbf{E}[K(s, \hat{s}_m)]\} + C(\phi, \Gamma, \rho_{min}, \rho_{max}, \beta, \Sigma).$$

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Decomposition

As in [3], we write:

$$K(s, \hat{s}_{\hat{m}}) \leq K(s, \bar{s}_m) + \bar{\gamma}(\bar{s}_m) - \bar{\gamma}(\hat{s}_{\hat{m}}) - \text{pen}(\hat{m}) + \text{pen}(m),$$

with $\bar{\gamma}(u) = \gamma(u) - \mathbf{E}[\gamma(u)]$ and $\bar{s}_m = \arg \min_{S_m} K(s, u)$.

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The χ^2 statistic

(1) $= (\bar{\gamma}(\bar{s}_{m'}) - \bar{\gamma}(\hat{s}_{m'})) \rightarrow$ introduce chi-square statistic :

$$\chi_m^2 = \chi^2(\bar{s}_m, \hat{s}_m) = \sum_{J \in m} Z_J, \text{ with } Z_J = \frac{[Y_J - E_J]^2}{E_J}.$$

$$Y_J = \sum_{t \in J} Y_t, \quad E_J = \mathbf{E}[Y_J]$$

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To apply Bernstein's inequality [6] we need:

- A space where Z_J is easily controlled
- The control of Y_J around its expectation.

Applying Bernstein's inequality

We define

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and $P[\chi_m^2 \mathbf{1}_{\Omega_m} \geq |m| + 8(1+\varepsilon)\sqrt{x|m|} + 4(1+\varepsilon)x] \leq e^{-x}$.

Segmentation procedure

$$\hat{K} = \arg \min_k \left\{ \gamma(\hat{s}_k) + \beta k \left(1 + 4 \sqrt{1.1 + \log \frac{n}{k}} \right)^2 \right\}$$

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- ④ tune β using the slope heuristic[1]

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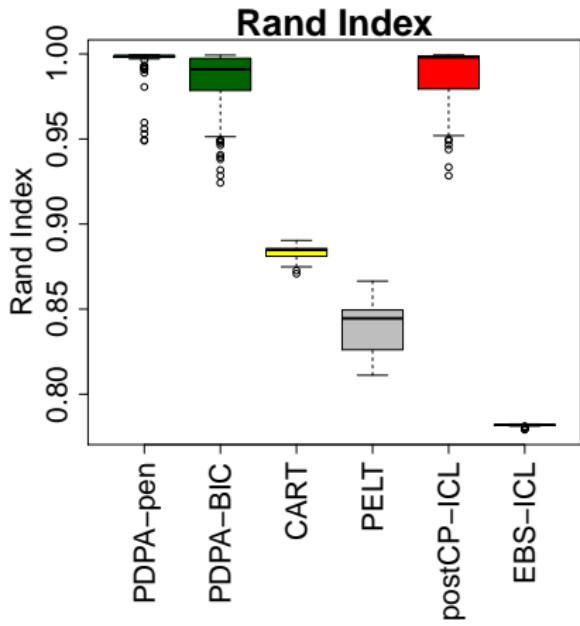
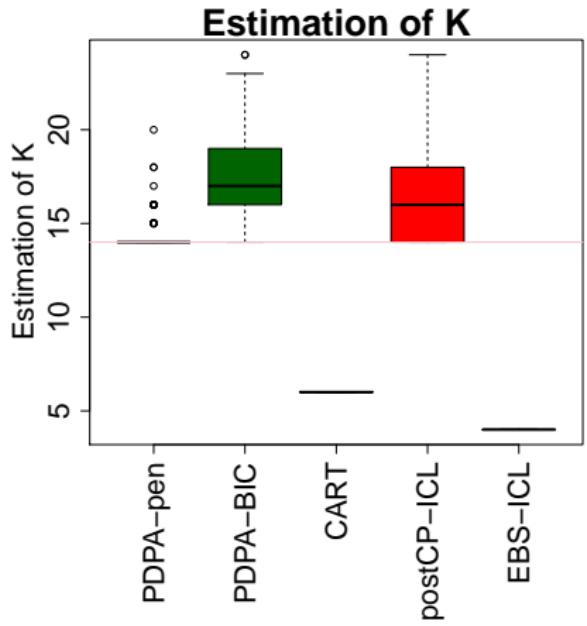
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Short-signal analysis

Resampling from a Drosophila gene with two isoforms
 $n = 5000$; $K = 14$

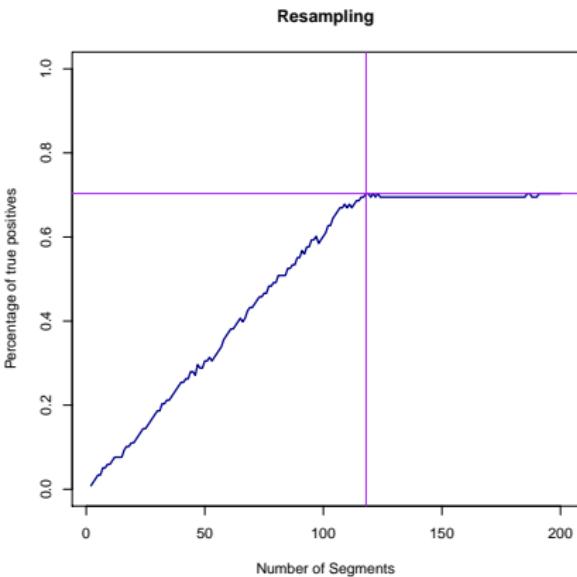
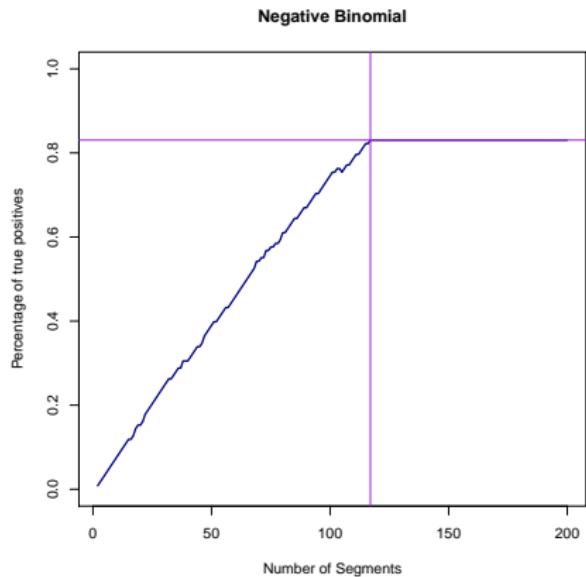


Long-signal analysis

Two simulation studies: $n = 230000$; $K = 118$

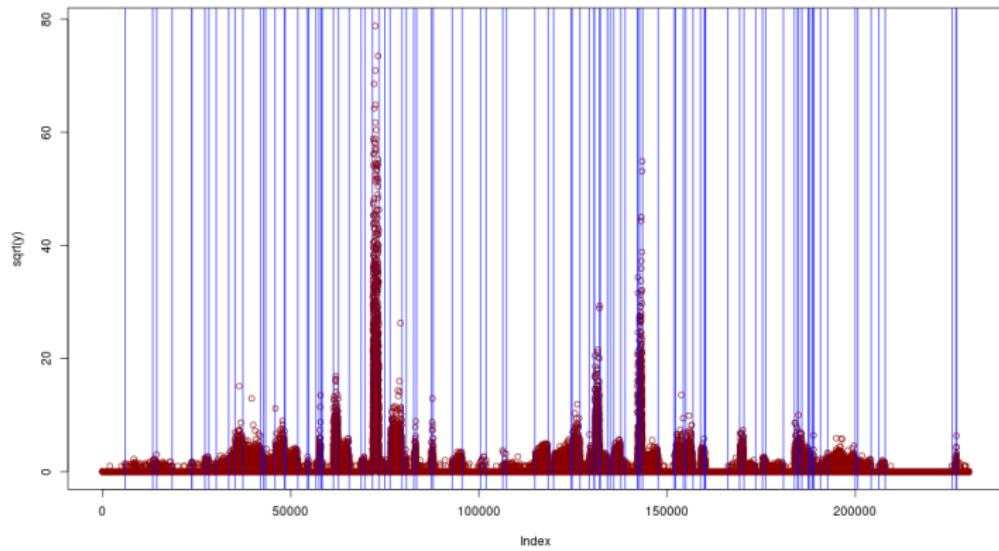
S1= simulation from negative binomial with 4 levels

S2= simulation by resampling real RNA-seq data pooled in 4 categories



Application to real data

Real RNA-Seq data: chromosome 1 from yeast genome.
 $\hat{K} = 103$



Conclusion

- Proposition of an estimator for the overdispersion ϕ :
→ inspired from [5], based on moment estimator
- Assuming ϕ is known:
 - Effective procedure for the segmentation of RNA-Seq data,
→ complexity in $\mathcal{O}(n \log(n))$
 - Theoretical guarantees,
→ oracle inequality
 - Excellent results in practice
- Complete procedure available in R package Segmentor3IsBack on the CRAN.

Thank you!

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